Comparison Between the Original and Shortened Versions of the National Institutes of Health Stroke Scale in Ischemic Stroke Patients of Intermediate Severity

Chun Fan Lee, PhD; Narayanaswamy Venketasubramanian, FRCP; K.S. Lawrence Wong, MD; Christopher L.H. Chen, FRCP; for the CHIMES Study Investigators

Background and Purpose—The 15-item National Institutes of Health Stroke Scale (NIHSS) has been critiqued for its complexity and variability, and shortened versions have been proposed. This study aimed to compare the measurement properties of the original version with 3 shortened versions with 11, 8, and 5 items, respectively.

Methods—Analyses were performed using data from an international, double-blind randomized controlled trial investigating the efficacy of MLC601 on stroke recovery in patients with ischemic stroke of intermediate severity (Chinese Medicine Neuroaid Efficacy on Stroke recovery [CHIMES]). To compare discriminative ability and responsiveness to change, the effect sizes of the NIHSS scores in relation to modified Rankin Scale, mini-mental status examination, and Barthel index were estimated using regression analysis.

Results—For both discriminative ability and responsiveness to change, the original version exhibited a larger effect size (0.55 and 0.84) in relation to modified Rankin Scale than the other 3 shortened versions (0.35–0.46 and 0.74–0.78).

Conclusions—The original 15-item NIHSS retained information that made it more discriminative and responsive to change than the shortened versions. We recommend future clinical researchers to use the full version NIHSS to evaluate patients’ stroke severity.

Clinical Trial Registration Information—URL: http://www.clinicaltrials.gov. Unique identifier: NCT00554723.

(Stroke. 2016;47:236-239. DOI: 10.1161/STROKEAHA.115.011657.)

Key Words: Barthel index ☐ MLC601 ☐ modified Rankin scale ☐ National Institute of Health Stroke Scale ☐ NeuroAiD ☐ outcome ☐ stroke severity

The National Institutes of Health Stroke Scale (NIHSS) is a widely used scale in assessing neurological deficits in stroke patients. Its measurement properties and accuracy in predicting clinical outcomes have also been extensively evaluated.1–4 Although it contains only 15 items and takes <8 minutes to perform,5 it has been critiqued for its complexity and variability.6–8 As a result, shortened versions of the NIHSS have been suggested.7–9 Recently, we conducted a large-scale clinical trial that used the NIHSS, as well as other outcome measures commonly used in ischemic stroke patients. Using the data from this trial, the objective of the current study was to compare the measurement properties of the original NIHSS with the shortened versions. Hereafter, the original version is referred to as NIHSS-15, and the 3 shortened versions, 11-item, 8-item, and 5-item, are referred to as NIHSS-11, NIHSS-8, and NIHSS-5, respectively. The 1-item version, however, was not included in the comparison.

Methods

Analyses were performed using data from the Chinese Medicine Neuroaid Efficacy on Stroke recovery (CHIMES) study, an international, randomized, placebo-controlled, double-blind trial investigating the efficacy of MLC601 on stroke recovery in patients with ischemic stroke of intermediate severity in the preceding 72 hours (clinicaltrials.gov NCT00554723).10,11 Patients were randomly allocated to either MLC601 or placebo for 3 months as add-on to standard stroke care.

Patients were assessed at baseline, day 10 (±2 days), or discharge, whichever was earlier, and month 3 (±1 week). At baseline, demographic and clinical information, as well as prestroke modified Rankin Scale (mRS) score, were ascertained. NIHSS and mini-mental status examination (MMSE) were also performed. At day 10 (or discharge, if earlier) and month 3 visits, NIHSS, mRS, and MMSE were assessed again. In addition, month 3 assessment included the Barthel index (BI).

Received September 28, 2015; final revision received September 28, 2015; accepted October 13, 2015.
From the Department of Biostatistics, Singapore Clinical Research Institute, Singapore (C.F.L.); Raffles Neuroscience Centre, Raffles Hospital, Singapore (N.V.); Department of Medicine and Therapeutics, Chinese University of Hong Kong, Prince of Wales Hospital, Shatin, New Territories (K.S.L.W.); and Department of Pharmacology, National University of Singapore, Clinical Research Centre, Singapore (C.L.H.C.).
Guest Editor for this article was Tatjana Rundek, MD, PhD.
The online-only Data Supplement is available with this article at http://stroke.ahajournals.org/lookup/suppl/doi:10.1161/STROKEAHA.115.011657/-/DC1.
Correspondence to Chun Fan Lee, PhD, Singapore Clinical Research Institute, 31 Biopolis Way, Nanos No 02-01, Singapore 138669. E-mail chunfanlee8128@gmail.com
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Stroke is available at http://stroke.ahajournals.org DOI: 10.1161/STROKEAHA.115.011657

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Subjects recruited for this trial were ischemic stroke patients having a prestroke mRS of 0 or 1 and baseline NIHSS-15 score between 6 and 14; hence, the baseline measures did not vary sufficiently for the evaluation of measurement validity. Therefore, in this study, the data obtained at day 10 and month 3 were analyzed. Patients who died before the assessment at month 3 were excluded from the analysis. Main analyses were based on all eligible patients, whereas subgroup analyses by treatment were also conducted as sensitivity analyses to examine the robustness of the results.

The discriminative ability and responsiveness to change of the 4 NIHSS scores were compared by regressing each score on other

<table>
<thead>
<tr>
<th>NIHSS version</th>
<th>Model $R^2$</th>
<th>mRS</th>
<th>Effect Size</th>
<th>Difference (95% CI)</th>
<th>MMSE</th>
<th>Effect Size</th>
<th>Difference (95% CI)</th>
<th>Barthel Index (per 5 Points)</th>
<th>Effect Size</th>
<th>Difference (95% CI)</th>
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<tbody>
<tr>
<td>Discriminative ability</td>
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<tr>
<td>NIHSS-15</td>
<td>0.64</td>
<td>0.55</td>
<td>0.00 (0.00–0.13)</td>
<td>0.05</td>
<td>0.00 (0.00–0.04)</td>
<td>0.00 (0.00–0.02)</td>
<td>0.05</td>
<td>0.00 (0.00–0.02)</td>
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<tr>
<td>NIHSS-11</td>
<td>0.62</td>
<td>0.46</td>
<td>0.09 (0.04–0.13)</td>
<td>0.04</td>
<td>0.01 (0.01–0.02)</td>
<td>0.01 (0.01–0.02)</td>
<td>0.04</td>
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<tr>
<td>NIHSS-8</td>
<td>0.59</td>
<td>0.46</td>
<td>0.09 (0.04–0.13)</td>
<td>0.05</td>
<td>0.01 (0.01–0.02)</td>
<td>0.01 (0.01–0.02)</td>
<td>0.05</td>
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<tr>
<td>NIHSS-5</td>
<td>0.57</td>
<td>0.35</td>
<td>0.20 (0.12–0.27)</td>
<td>0.05</td>
<td>0.00 (0.00–0.02)</td>
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<td>Responsiveness to change</td>
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<tr>
<td>NIHSS-15</td>
<td>0.57</td>
<td>0.84</td>
<td>0.00 (0.00–0.13)</td>
<td>0.06</td>
<td>0.00 (0.00–0.04)</td>
<td>0.00 (0.00–0.02)</td>
<td>0.06</td>
<td>0.00 (0.00–0.02)</td>
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<tr>
<td>NIHSS-11</td>
<td>0.55</td>
<td>0.74</td>
<td>0.10 (0.06–0.15)</td>
<td>0.05</td>
<td>0.01 (0.01–0.02)</td>
<td>0.01 (0.01–0.02)</td>
<td>0.05</td>
<td>0.01 (0.01–0.02)</td>
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<tr>
<td>NIHSS-8</td>
<td>0.47</td>
<td>0.78</td>
<td>0.06 (0.00–0.10)</td>
<td>0.07</td>
<td>0.00 (0.00–0.02)</td>
<td>0.01 (0.01–0.02)</td>
<td>0.07</td>
<td>0.00 (0.00–0.02)</td>
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<tr>
<td>NIHSS-5</td>
<td>0.40</td>
<td>0.77</td>
<td>0.07 (0.01–0.12)</td>
<td>0.05</td>
<td>0.01 (0.01–0.02)</td>
<td>0.01 (0.01–0.02)</td>
<td>0.05</td>
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</table>

BI indicates Barthel Index; CI, confidence interval; MMSE, mini-mental status examination; mRS, modified Rankin Scale; NIHSS, National Institutes of Health Stroke Scale; NIHSS-5, 5-item NIHSS; NIHSS-8, 8-item NIHSS; NIHSS-11, 11-item NIHSS; and NIHSS-15, original 15-item NIHSS.
stroke measures in a linear regression model. For discriminative ability, the NIHSS scores at month 3 were regressed on mRS and MMSE, and BI divided by 5 at month 3, adjusting for baseline characteristics and risk factors. The resulting regression coefficients can be interpreted as the change in NIHSS score per unit increment in mRS and MMSE and every 5-point increment in BI, respectively. Because the 4 NIHSS scores are not of the same metric and not directly comparable, standardization by dividing these coefficients by the residual standard deviation of the regression model was performed to obtain the effect size. This effect size may be regarded as a signal-to-noise ratio reflecting measurement precision: the larger the signal (numerator) and the smaller the noise (denominator), the larger the effect size. A score with a larger effect size is desired because it requires a smaller sample size to achieve the same research purpose. The 95% confidence interval for the difference between the effect sizes of the original and shortened versions was estimated using 1000 bootstrap replications. Similarly, responsiveness to change from day 10 to month 3 was also compared. Confirmatory factor analysis and receiver operating characteristic curve analysis were also performed (please see the online-only Data Supplement). All analyses were performed using SAS v9.3.

Results
The overall baseline characteristics and study flow of patients in the CHIMES study were previously described. In this study, we used the data from 1071 patients at day 10 and 949 patients at month 3 who were alive and had complete NIHSS assessments. Their mRS, MMSE, BI, and NIHSS scores are summarized in Table 1. There was no significant difference in these scores between the 2 groups.

The upper panel of Table 2 shows the effect size of the 4 NIHSS scores for detecting a difference in mRS, MMSE, and BI at month 3 adjusting for baseline characteristics and risk factors, together with the bootstrap confidence interval for the difference with NIHSS-15. For mRS, the original NIHSS-15 score had a significantly larger effect size (0.55) than the other 3 shortened versions (0.35–0.46). However, the effect sizes were similar among the 4 versions and of small magnitude for MMSE (−0.05 to −0.04) and BI (−0.17 to −0.14). Similarly, the lower panel of Table 2 compares the effect size of the NIHSS scores for detecting a change from day 10 to month 3 in mRS and MMSE. Again, NIHSS-15 achieved a larger effect size (0.84) for mRS than the other 3 versions (0.74–0.78), whereas the 4 versions had similar effect size for MMSE (−0.06 to −0.06).

A set of sensitivity analyses stratified by treatment group were also performed for the above analyses. Results were not qualitatively different from that based on all patients and, thus, are not repeated here. Results for confirmatory factor analysis and receiver operating characteristic curve analysis were summarized in online-only Data Supplement (Table I, Figure I, and Figure II in the online-only Data Supplement).

Discussion
Our comparison of the original version with the shortened versions of the NIHSS showed that the original 15-item NIHSS exhibited the largest effect size in discriminative ability and responsiveness to change in relation to mRS. However, regardless of the version used, the NIHSS scores seemed to be only weakly to moderately associated with the MMSE and BI, suggesting that these instruments may measure domains and functions not addressed by the NIHSS and, hence, function as complementary outcomes.

Our study does have clinical and research implications. Busy clinicians are often tempted by shortened versions of clinical scales to allow more efficient use of their time with the patient. Researchers would prefer fewer data points as this reduces data entry time and occasions for errors. However, the shortened NIHSS, although easier and quicker to perform than the full version, comes at a cost of lowered discriminatory value and responsiveness to change, which may impact negatively on the ability to predict outcome, be it in clinical or research settings.

A limitation of this study was the relatively stringent inclusion criteria of the trial, which recruited patients with narrow ranges of prestroke mRS and baseline (original version) NIHSS. Hence, our findings may not be extrapolated beyond the population included in the study, that is, more severe strokes and patients dependent before the index stroke. The shortened versions may be more discriminative and useful in this group of more severe stroke patients. The strengths of this study include that it is a multicentre study performed by experienced stroke trialists, which would provide reliable data. The study involved a large number of subjects, the vast majority of whom had complete NIHSS and functional outcome data.

In summary, despite being critiqued for its complexity and variability, the original 15-item NIHSS, which can be performed easily over a few minutes, retained information that made it more discriminative and responsive to change than the shortened versions. We recommend that clinical researchers use the full version NIHSS to evaluate patients’ stroke severity.

Appendix—CHIMES Investigators

Sources of Funding
The CHIMES study was supported by grants from the National Medical Research Council of Singapore (NMRC/11288/2011 and NMRC/1096/2006) and the CHIMES Society. The authors received funding for the trial and accommodation and transportation support for meetings from the CHIMES Society. Moleac (Singapore) provided grants to the CHIMES Society of which the society had sole discretion on use.

Disclosures
None.

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*Stroke*. 2016;47:236-239; originally published online December 1, 2015;
doi: 10.1161/STROKEAHA.115.011657

*Stroke* is published by the American Heart Association, 7272 Greenville Avenue, Dallas, TX 75231
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Print ISSN: 0039-2499. Online ISSN: 1524-4628

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